

Emerging Market Business Cycles with Remittance Fluctuations

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This paper analyzes the implications of remittance fluctuations for various macroeconomic variables and sudden stops. The paper employs a quantitative two-sector model of a small open economy with financial frictions calibrated to Mexican and Turkish economies, two major recipients, whose remittance receipts feature opposite cyclical characteristics. We find that remittances dampen business cycles in Mexico, whereas they amplify the cycles in Turkey. Their quantitative effects in the long run, approximated by the stochastic steady state, are mild. In the short run, however, remittances have quantitatively large impacts on the economy, when the economy is borrowing-constrained. This is because agents in the economy cannot adjust their precautionary wealth to sudden tightening in credit, and hence, fluctuations in remittances get magnified through an endogenous debt-deflation mechanism. The findings suggest that procyclical (or countercyclical) remittances can play a significant deepening (or mitigating) role for sudden stops. [JEL F41, F32, E32]

IMF Staff Papers (2010) **57**, 303–325. doi:10.1057/imfsp.2009.3;
published online 15 December 2009

Officially recorded migrant remittances received by developing countries increased to \$250 billion in 2007, representing a 50 percent increase in just three years since 2005 by the World Bank

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estimates.¹ Thanks to this fast growth, the total amount officially received by the developing world has more than tripled in nominal terms since the beginning of this decade. This growth has been visibly faster than the growth of private capital flows and official development assistance (ODA), enabling remittances to eventually surpass nonforeign direct investment (FDI) (private debt and portfolio equity) and ODA flows, and to almost catch FDI receipts in magnitude as of 2004.² As a result, remittances have become a more important source of foreign exchange than private capital flows, ODA, and even FDI for many developing countries. Popular stance in the policy circles and the academic literature is to view this rapid growth as a generally positive development for developing economies on account of the following:

1. In contrast to other capital flows, remittances do not create any liabilities such as debt servicing or profit transfers in the future.
2. Remittance flows are usually more stable than private capital flows including FDI (Ratha, 2003; Buch and Kuckulenz, 2004).
3. Remittances could serve as macroeconomic stabilizers, as it is often argued, because migrant workers are likely to increase the amounts transferred to help family members left behind, whenever the economic activity back home slows down (UNCTAD, 2006; World Bank, 2006a and 2006b).
4. Remittance receipts may promote entrepreneurship, investment in physical capital and human capital formation, by helping relax borrowing constraints facing family members that stayed home (Yang, 2008).

Yet, whether high remittance receipts are always a blessing depends on the nature of comovements, if any, between business cycles in the home countries of migrants and cyclical fluctuations in the remittance flows. Remittances will move countercyclically to output fluctuations in the home countries of migrant workers, if the dominant motivation behind their remitting decisions is to contribute to the financing of consumption expenditures of family members left behind (the so-called altruistic consumption-smoothing motive).³ However, an increasing number of studies after Sayan (2004) have pointed out the possibility of procyclical remittances due to investment

¹By the definition in the World Bank's *Global Development Finance 2003*, migrant remittances are made up of workers' remittances, compensation of employees, and migrants' transfers.

²During the same year, remittance receipts exceeded combined public and private capital inflows in 36 developing countries and were larger than total merchandise exports in 12 others. In some countries such as Mexico, FDI receipts often fall short of remittances (World Bank, 2006a).

³In this case, household members working abroad would increase the amounts they remit when there is a recession/crisis in the home economy so as to help compensate the decline in household income due to unemployment and wage cuts that family members may face during such episodes.

or portfolio-diversification motive—see Lueth and Ruiz-Arranz (2007) and papers cited therein.⁴

Regardless of the underlying motivation to remit, remittances could be a blessing, if they move counter to home country business cycles, as they will then serve as macroeconomic stabilizers against cyclical contractions or sudden stops (see Bugamelli and Paternò, 2005). If they are procyclical, on the other hand, they could be a setback, as the drops in remittance receipts observed during cyclical contractions or sudden stops would magnify the damage resulting from such contractions or stops. Answers to how effective countercyclical or procyclical remittance flows could be in lowering or increasing the amplitude of macroeconomic fluctuations depend on several factors and are less obvious. In addition to the nature of comovements between remittance fluctuations and business cycles, the response time of remittances to business cycle movements, and the share of remittances in gross domestic product (GDP) need to be taken into consideration while answering this question. Likewise, the quantitative effects that remittance fluctuations could have on different macroeconomic variables during sudden stops experienced by the recipient economies need to be investigated quantitatively, using an appropriate model that captures general equilibrium interactions between key macroeconomic variables.

We aim to shed light on these issues by examining the effect of remittance flows with opposite cyclical characteristics on different macroeconomic aggregates and sudden stops.⁵ For this purpose, we introduce remittances into the small open economy model of Mendoza (2005) and calibrate it to the data for the Mexican and Turkish economies, two major recipients which differ with regard to the way remittance receipts respond to respective home country business cycles. Remittances are procyclical in Turkey, whereas they are countercyclical in Mexico, as thoroughly discussed by Sayan and Tekin-Koru (2008a).

The model features a tradable sector and a nontradable sector in which the liabilities are denominated in units of tradable goods (that is, liabilities are dollarized), and agents face a borrowing constraint in international capital markets. Foreign debt is partially leveraged through income generated in the nontradable sector. Interaction of these two frictions, that is, liability dollarization and the borrowing constraint, creates a debt-deflation mechanism that mimics the key features of sudden stops experienced by both Mexico and Turkey. In the absence of remittances, a

⁴Sayan (2006) presented evidence showing that remittances are procyclical in some countries and listed a number of possible reasons underlying this procyclicality. Our results also confirm that cyclical characteristics of remittances may be different across countries (see Figures 1 and 2, and Table 1, which are generated using data in constant local currency units, that is, real terms).

⁵In a cross-country study, Bugamelli and Paternò (2005) find that remittances, as cheap inflows of foreign currencies, might reduce the probability that foreign investors suddenly flee out of emerging and developing economies, triggering a dramatic current account adjustment.

shock to the economy making the borrowing constraint binding leads to a decline in tradable consumption and relative price of nontradables. The decline in relative price of nontradables tightens the constraint even further because the collateral value of nontradable income becomes lower. Further tightening of the borrowing constraint creates a feedback mechanism, which eventually leads to a collapse in consumption and relative price of nontradables and reversals in current account.⁶ Using this model, we quantitatively explore how important the remittance fluctuations are to countries where they move countercyclically and procyclically to domestic business cycles, and investigate possible effects of these fluctuations on sudden stops experienced by such countries.

Our results indicate that remittances dampen the business cycles in the Mexican economy, whereas they deepen the cycles in the Turkish economy as expected. Their quantitative effects in the long run approximated by the stochastic steady state are rather mild, and do not significantly depend on whether the economy is borrowing constrained or not.⁷ In the short run, however, remittances can have quantitatively large impacts on the macroeconomy, if the borrowing constraints are binding. In the short run, agents in the economy cannot adjust their precautionary wealth to sudden tightening of credit, causing small remittance shocks to the economy to get magnified through the endogenous debt-deflation mechanism.

We quantify the short-run impact effects of remittance fluctuations using forecasting functions.⁸ We compare the impact effect of income shocks with and without the accompanying remittance shocks. In the Turkish case, a one-standard-deviation negative remittance shock that accompanies a one-standard-deviation income shock magnifies the decline in tradable consumption by 2 percent and the reversal in current account-GDP ratio by 3 percentage points. In the Mexican case, a one-standard-deviation positive remittance shock that follows the negative income shock smoothes the decline in tradable consumption by 1.4 percent and decreases the reversal in the current account-GDP ratio by 2 percentage points. These results suggest that remittances can have significant amplifying, in the case of procyclical remittances, or smoothing, in the case of countercyclical remittances, effects on sudden stops.

⁶Various studies (including Calvo, Izquierdo, and Mejía, 2004; Mendoza, 2002) showed that such credit crunches amplified by highly dollarized liabilities were the main driving force of the sudden stops that emerging markets like Mexico and Turkey faced during the last decade and a half. It is merely this mechanism in the model that generates sudden-stop-like crises dynamics.

⁷This result mimics the findings of Mendoza (2002), who finds that imposition of borrowing constraints do not alter the long-run business cycles quantitatively, because agents engage in precautionary savings and minimize the impacts of borrowing constraints on the macroeconomy in the long run.

⁸Forecast functions are a variant of impulse response functions, which are derived by setting the initial conditions of the economy to a state where the economy is prone to a sudden stop.

Despite their significance, the current literature lacks studies on such effects of remittances except through cross-country regressions (see Bugamelli and Paternò, 2006).⁹ Furthermore, the existing studies largely focus on the effects of countercyclical remittances on the volatility of output, consumption, and investment in the recipient countries (see IMF, 2005), and overlook the macroeconomic effects of procyclical remittance flows that individual countries such as Turkey may receive. This paper provides a first look at these issues using a dynamic stochastic general equilibrium model with sudden stops.¹⁰ In addition to our work, there are a limited number of studies that use a stochastic general equilibrium framework to explore the effects of remittances on different macroeconomic variables in small open economy settings, but none of them focuses on sudden stops. Chami, Cosimano, and Gapen (2006) and Jansen, Naufal, and Vacaflares (2007) investigate the effects of remittances on key macroeconomic variables, and the conduct of fiscal/monetary policies by using the same countercyclical specification which relates changes in remittances to output so that they would increase when there is a downturn in the recipient economy. Acosta, Lartey, and Mandelman (2007) consider the effects of altruistically motivated (countercyclical) and investment oriented (procyclical) remittances separately but in contrast to the present study, they focus on the effects of remittances through their potential to cause Dutch disease.

Prior to these studies, Chami, Fullenkamp, and Jahjah (2003) used a model with micro foundations to formalize the argument about the altruism of migrants as the underlying reason for the countercyclicity of remittances, and backed this theoretical result with panel data evidence indicating that remittances respond negatively to changes in output. Later, other studies such as IMF (2005); Mishra (2005); and World Bank (2006a and 2006b) presented additional evidence indicating a negative relationship between output and remittance receipts. Yet, the first look in the literature by Sayan (2004) at the comovements between the cyclical components (defined as deviations from trend) of home country output and remittances series produced different results. Using quarterly time series data on remittances sent home by Turkish workers in Germany, Sayan (2004) found that remittance receipts of Turkey from Germany were procyclical.¹¹

⁹See Aguinias (2006) for an extensive review of this literature.

¹⁰In a recent cross-country study, Bugamelli and Paternò (2005) find that as cheap inflows of foreign currencies, remittances might reduce the probability that foreign investors suddenly flee out of emerging and developing economies and trigger a dramatic current account adjustment.

¹¹The sharp drop in remittance receipts to Turkey during the 2001 crisis conformed to these findings (Ratha, 2003). Using a larger country sample, Sayan (2006) provided additional support to the view that nonaltruistic considerations may drive the remittance behavior, leading to procyclical remittances. These findings are in line with the intuition that Buch, Kuckulenz, and Le Manchec (2002) and Ratha (2003) previously pointed out.

More recently, Sayan and Tekin-Koru (2008a) considered the cyclical behavior of Turkish remittances from Germany and Mexican remittances from the United States in a comparative study. Using quarterly data covering the 1980s onward, they found that remittance receipts of Mexico from the United States were synchronously countercyclical to the business cycle in Mexico, whereas Turkish remittances were again procyclical and followed the business cycle in Turkey with a one-quarter lag.¹²

In addition to the literature on remittances, our work is also related to the literature on business cycle fluctuations in small open economies as exemplified by the works of Durdu, Mendoza, and Terrones (forthcoming); Durdu (2009); Durdu and Mendoza (2006); Neumeyer and Perri (2005); Kose (2002); and Mendoza (1991 and 2002 and 2005). Mendoza (1991) provides a workhorse quantitative small open economy model that accounts for the aggregate fluctuations in small open economies. Kose (2002) extends this model to explore the importance of world price shocks and fluctuations in world interest rates on business cycles of small open economies. Neumeyer and Perri (2005) focus on the importance of world interest rates in driving business cycles. Our paper relates to this first group of papers by shedding some light on the business cycle implications of remittance fluctuations. Mendoza (2002 and 2005), among others, emphasizes the role of frictions in the world capital markets and accounts for the observed features of sudden stops. Durdu (2009) examines how hedging and self-insurance options and their implications for sudden stops are affected if the agents have access to GDP-indexed credit contracts. This paper contributes to this literature by examining the quantitative importance of remittance fluctuations on sudden stops by employing the features used in those studies.

I. Model

We introduce remittances to the two-sector small open economy model of Mendoza (2005). Foreign debt is denominated in units of tradables and imperfect credit markets impose a borrowing constraint that limits external debt to a share of the value of total income in units of tradables.

Representative households receive stochastic exogenous remittances denoted $(1 + \varepsilon_t^R)Rem$; a stochastic endowment of tradables and a non-stochastic endowment of nontradables, which are denoted $(1 + \varepsilon_t^Y)Y^T$ and y^N , respectively. ε_t^R and ε_t^Y are respective shocks to the remittances and tradables endowments. Households derive utility from aggregate consumption (c), and

¹²Sayan and Tekin-Koru (2008b) offered the stagnation of Turkish migration to Germany in the 1980s after family reunifications as the most plausible explanation for the procyclicality of remittances. Turkish migrants' ties with the family members remaining in Turkey weakened with the passage of time, they argued, causing altruism motive to lose its strength. For Mexico, on the other hand, this motive must remain strong as migration of Mexican workers to the United States continues, albeit mostly illegally.

they maximize the following stationary cardinal utility function:

$$U = E_0 \left\{ \sum_{t=0}^{\infty} \exp \left[- \sum_{t=0}^{t-1} \gamma \log(1 + c_t) \right] u(c_t) \right\}. \quad (1)$$

Functional forms are given by

$$u(c_t) = \frac{c_t^{1-\sigma} - 1}{1 - \sigma}, \quad (2)$$

$$c_t(c_t^T, c_t^N) = [\omega(c_t^T)^{-\mu} + (1 - \omega)(c_t^N)^{-\mu}]^{-\frac{1}{\mu}}. \quad (3)$$

The instantaneous utility function is in constant relative risk aversion (CRRA) form with an intertemporal elasticity of substitution value of $1/\sigma$. Aggregate consumption is given by a constant elasticity of substitution (CES) function, where $1/(1 + \mu)$ is the elasticity of substitution between consumption of tradables and nontradables and where ω is the CES weighing factor. $\exp[-\sum_{t=0}^{t-1} \gamma \log(1 + c_t)]$ is an endogenous discount factor that is introduced to induce stationarity in consumption and asset dynamics. γ is the elasticity of the subjective discount factor with respect to consumption.¹³

The households' budget constraint is

$$c_t^T + p_t^N c_t^N = (1 + \varepsilon_t^R) Rem + (1 + \varepsilon_t^y) y^T + p_t^N y^N - b_{t+1} + (1 + r) b_t, \quad (4)$$

where b_t is current bond holdings, $(1 + r)$ is the gross return on bonds, and P_t^N is relative price of nontradables. Notice that bond returns are denominated in units of tradables, whereas they are partially financed by income earned in nontradable sector, that is, liabilities are dollarized.

In addition to the budget constraint, foreign creditors impose the following borrowing constraint, which limits debt issuance as a share of total income at period t not to exceed κ . Moreover, the bond holdings cannot be lower than a minimum level, Ω :¹⁴

$$b_{t+1} \geq -\kappa[(1 + \varepsilon_t^R) Rem + (1 + \varepsilon_t^y) y^T + p_t^N y^N] \geq \Omega. \quad (5)$$

The borrowing constraint takes a similar form to those used in the sudden stops literature (see Mendoza, 2005; Caballero and Panageas, 2003). The interaction of the borrowing constraint with the liability dollarization

¹³Mendoza (1991) first introduced preferences with endogenous discounting to quantitative small open economy models. See Schmitt-Grohé and Uribe (2003); and Kim and Kose (2003) for alternative specifications used for this purpose.

¹⁴This lower bound for bond holdings (or upper bound for debt level) is introduced to rule out equilibria in which the constraint is satisfied at very high levels of debt that increase c^T and p^N . See Mendoza (2005) for further details.

induces a debt-deflation mechanism that amplifies the shocks to the economy (see Mendoza (2005) for further details.)

The optimality conditions can be summarized as follows:

$$U_c(t) \left(1 - \frac{v_t}{\lambda_t}\right) = \exp[-\gamma \log(1 + c_t)] E_t \left\{ \frac{(1+r)p_t^c}{p_{t+1}^c} U_c(t+1) \right\}, \quad (6)$$

$$\frac{1 - \omega}{\omega} \left(\frac{c_t^T}{c_t^N}\right)^{1+\mu} = p_t^N, \quad (7)$$

along with the budget constraint (equation (4)), the borrowing constraint (equation (5)), and the standard Kuhn-Tucker conditions. v and λ are the Lagrange multipliers of the borrowing constraint and the budget constraint, respectively. U_c is the derivative of lifetime utility with respect to aggregate consumption. p_t^c is the CES price index of aggregate consumption in units of tradable consumption, which equals $\left[\omega^{\frac{1}{\mu}} + (1-\omega)^{\frac{1}{\mu}}(p^N)^{\frac{\mu}{1+\mu}}\right]^{\frac{1+\mu}{\mu}}$. Equation (6) is the standard Euler equation equating marginal utility at time t to that at time $t+1$. Equation (7) equates the marginal rate of substitution between tradables consumption and nontradables consumption to the relative price of nontradables.

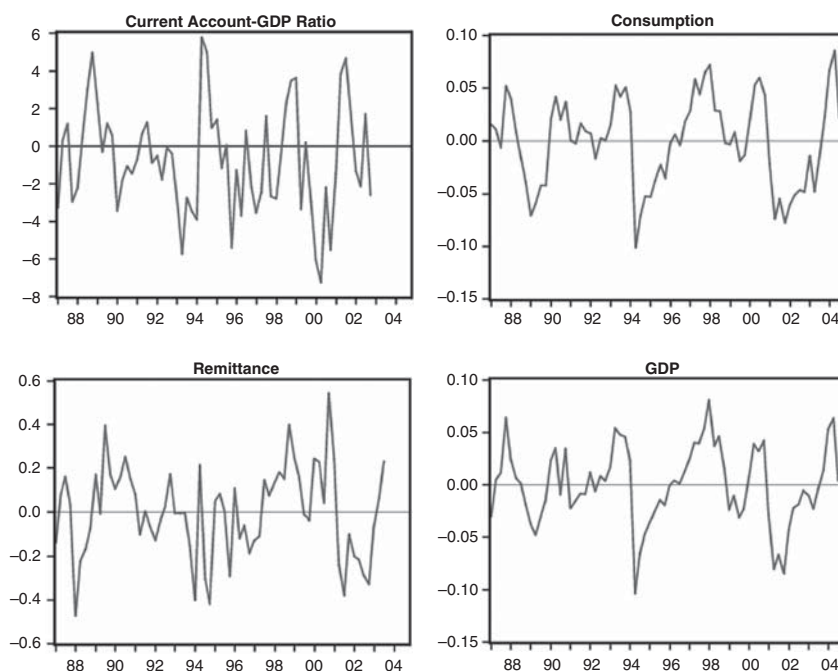
This endowment economy model is certainly not less powerful than a model with capital accumulation in terms of its capacity to explain sudden stop dynamics. One needs an amplification mechanism to generate sudden stop dynamics, and the amplification is generated through the interaction of the borrowing constraint with the relative price of nontradables in our setup. This amplification can alternatively be generated through the borrowing price of capital in a model with capital accumulation. Given that the magnitude of the amplification in this endowment economy is as high as the one with capital accumulation (see Durdu, Mendoza, and Terrones, forthcoming; Mendoza, 2002 and 2005), however, we opt for the simpler endowment economy setup.

Concerning the cyclical characteristics of remittances, we take the pro- or countercyclical nature of the remittances in Turkey and Mexico as given by the data (see Figures 1 and 2) and analyze their implications on macroeconomic aggregates, instead of delving into possible reasons and different motivations to remit underlying these characteristics. Sayan (2006), for example, points out that how much to remit is a complex decision involving many other factors than the migrants' altruistic desire to help family members smooth their consumption, and different variables driving remittance behavior might be differently affected by the state of economic activity over home country business cycles.¹⁵

The response of remittance flows to cyclical fluctuations in economic activity at home is indeed likely to be different, when remittances are primarily motivated by the differences in rates of return to savings in

¹⁵See Russell (1986) for a list of factors that may affect remittance flows.

Figure 1. Business Cycles in Turkey

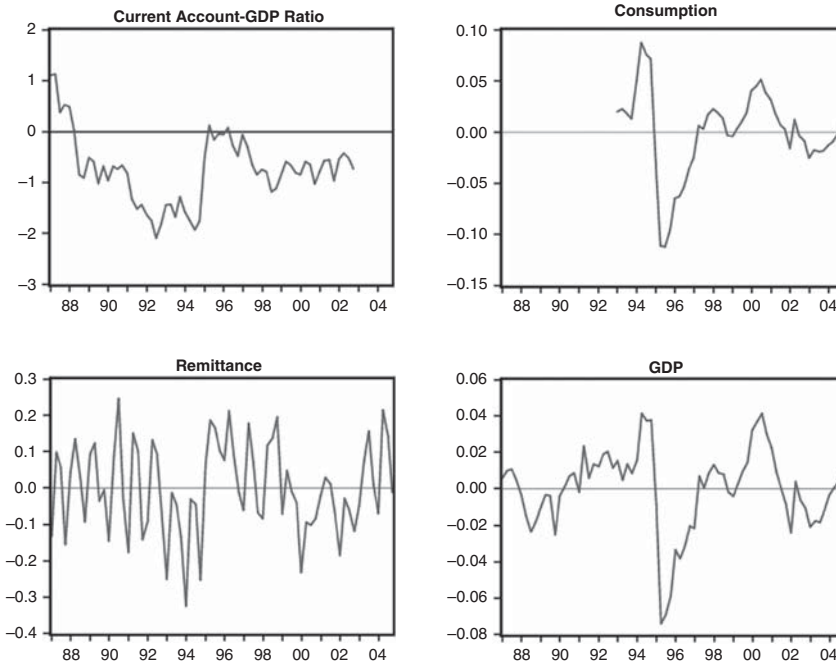


Sources: IMF, *International Financial Statistics*; Central Bank of Turkey.

Note: Data are quarterly seasonally adjusted real series in constant local currency units. Consumption, GDP, and remittance data are logged and filtered using an HP filter with a smoothing parameter value of 1,600. The figure shows deviations from trend for these variables over the sample period.

home and host countries of migrant workers (investment or portfolio-diversification motive). In such a case, upturns in economic activity may be associated with an increase in remittance receipts of the home country economy, whereas a downturn may lower these receipts, producing a procyclical remittance behavior. Thus, the response of remittance flows to cyclical output movements may differ, depending upon whether investment (or portfolio-diversification) motive is stronger than the altruistic consumption-smoothing motive for the migrant workers from different countries. Furthermore, which motive is stronger may change over time as the migrants' ties with relatives back home get weaker—due to increasing duration of stay in the host country and/or reunification of the migrants and immediate family members in the country of employment, and so on (Sayan and Tekin-Koru, 2008b). The passage of time may also allow remittance-receiving households to save enough to switch from wage earners to small entrepreneurs, possibly causing the remittance behavior of the family members abroad to change. Such a change in the labor force participation status of recipient households may also occur due to savings made possible by positive remittance shocks, because such remittance shocks often serve to

Figure 2. Business Cycles in Mexico



Sources: IMF, *International Financial Statistics*; Central Bank of Mexico.

Note: Data used in the figure are quarterly seasonally adjusted real series in constant local currency units. Consumption, GDP, and remittance data are logged and filtered using an HP filter with a smoothing parameter value of 1,600. The figure shows deviations from trend for these variables.

relax borrowing constraints facing households in developing countries as argued by Yang (2008).¹⁶

The next section presents the results of a series of numerical exercises that explore the implications of remittance fluctuations.

II. Quantitative Analysis

The recursive representation of the households' problem can be formulated as follows:

$$V(b, \varepsilon) = \max\{u(c) + (1 + c)^{-\gamma} E[V(b', \varepsilon')]\}$$

¹⁶As a matter of fact, remittance receipts of households from abroad often increase in domestic currency terms due to depreciation of local currency during recessions/crises, even if the amounts remitted stay the same in foreign currency terms. If the increase in remittance receipts in domestic currency more than compensates for the loss in household income, this may cause the labor force participation behavior of household members at home from wage earners to self-employed, as discussed by Yang (2008) within the context of the effects of the Asian crisis on the remittances from Filipino workers working abroad (see also Funkhouser, 1992).

s.t.

$$c^T = (1 + \varepsilon^R)Rem + (1 + \varepsilon^Y)y^T - b' + Rb, \quad (8)$$

$$c^N = y^N,$$

$$b' \geq -\kappa[(1 + \varepsilon^R)Rem + (1 + \varepsilon^Y)y^T + p^N y^N] \geq \Omega.$$

Here, $\mathbf{B} = \{b_1 < \dots < b_{NB}\}$ is the endogenous state space. $\varepsilon = \{\varepsilon^R, \varepsilon^Y\}$ is the exogenous state space, which follows a joint Markov process with known vectors of realization. To approximate the Markov process for those exogenous shocks, we first estimate a vector-autoregression (VAR) of tradable output and remittance series. Then, we estimate the Markov transition matrix using Tauchen and Hussey's (1991) quadrature procedure. The VAR representation of the system can be summarized as follows:

$$\xi_t = RHO.\xi_{t-1} + e_t, \quad (9)$$

where

$$\xi_t \equiv \begin{bmatrix} \varepsilon^Y \\ \varepsilon^R \end{bmatrix}, \quad RHO = \begin{bmatrix} \rho_y & \rho_{y,R} \\ \rho_{R,y} & \rho_R \end{bmatrix}, \quad e_t \equiv \begin{bmatrix} e_t^Y \\ e_t^R \end{bmatrix}.$$

We calibrate the model to both Turkish and Mexican economies. Statistics in Table 1 suggest that the remittance fluctuations are procyclical in Turkey, whereas they are countercyclical in Mexico. The parameter values are summarized in Table 2. Parameters common for both countries are relative risk aversion parameter, which is set to 2; world interest rate, which is set to the quarterly equivalent of 6.5 percent, and the elasticity of substitution between tradable and nontradable goods, which is set to 0.316 following the estimates of Ostry and Reinhart (1992). The relative price of nontradables and mean tradable endowments are normalized to 1 for both countries. The rest of the parameters are country specific as summarized in the table. The estimated VAR coefficients for Turkey are

$$RHO = \begin{bmatrix} 0.580 & 0.016 \\ 1.484 & 0.365 \end{bmatrix},$$

with the standard deviation of the endowment shock equal to 0.035 and of the remittance shock equal to 0.207. The VAR coefficients for Mexico are

$$RHO = \begin{bmatrix} 0.687 & -0.023 \\ -1.070 & 0.200 \end{bmatrix},$$

with the standard deviation of the endowment shock equals 0.027 and of the remittance shock equals 0.124.

We solve the stochastic simulations using value function iteration over a discrete state space. The state space spans $[-5.0, 3.0]$ interval with 1,000 grid points for both calibrations to Mexico and Turkey. We employ the solution procedure described in Durdu (2009) and Mendoza (2002). We start with an initial conjecture for the value-function and solve the model without imposing the borrowing constraint. We then check whether the bond decision

Table 1. Business Cycle Facts for Mexico and Turkey

Variables: x	$\sigma(x)$	$\sigma(x)/\sigma(Y)$	$\sigma(x)$	$\sigma(x, Y)$	Sudden Stop	Sudden Stop Relative to Standard
Mexico					1994:4–1995:1	
GDP (Y)	2.26	1.00	0.80	1.00	-7.40	3.30
Tradables GDP	2.70	1.20	0.70	0.92	-8.98	3.35
Nontradables GDP	2.19	0.97	0.83	0.98	-6.18	2.82
Remittance	12.38	5.48	0.26	-0.38	6.61	0.53
Consumption	4.22	1.87	0.84	0.97	-11.20	2.65
Real exchange rate	8.63	3.82	0.73	0.60	-32.84	3.81
CAY	0.70	0.31	0.83	-0.47	2.22	3.18
Turkey					1994:1–2	
GDP (Y)	3.70	1.00	0.67	1.00	-10.38	2.00
Tradables GDP	3.51	0.95	0.52	0.96	-10.93	3.11
Nontradables GDP	4.02	1.09	0.68	0.98	-10.01	2.49
Remittance	20.67	5.60	0.42	0.22	-30.36	1.47
Consumption	4.13	1.12	0.75	0.92	-10.10	2.40
Real exchange rate	9.11	2.47	0.68	0.60	-31.63	3.47
CAY	2.74	0.74	0.63	-0.59	9.70	3.38

Sources: Bank of Mexico; Central Bank of Turkey; IMF, *International Financial Statistics*.

Note: The data cover periods 1987:Q1–2004:Q4 for both Mexico and Turkey. Data are quarterly seasonally adjusted real series in constant local currency units. Consumption, GDP, and remittance data are logged and filtered using an HP filter with a smoothing parameter 1,600. Real exchange rates are calculated using the IMF definition ($RER_i = NER_i \times CPI_i / CPI_{US}$ for country i). CAY = current-account-GDP ratio.

satisfies the borrowing constraint. If so, the solution is found and we calculate the implied value-function, which is then used as a conjecture for the next iteration. If not, we impose the borrowing constraint with equality and solve it again. Then, we calculate the implied value-function using the optimal bond holdings and iterate to convergence.

We divide the stochastic simulations into four sets. In the first set, which we call *Baseline NB* (for “nonbinding economy”), the borrowing constraint does not bind and the economy is hit by both endowment and remittance shocks. In the second set, which is labeled *Baseline B*, the economy is hit by both endowment and remittance shocks again, but it, now, faces a borrowing constraint. In the third set, the economy is hit by an endowment shock only, and the borrowing constraint does not bind. This set is labeled *End. Shock NB* accordingly. In the last set, called *End. Shock B*, the economy is hit by endowment shock only, and the borrowing constraint binds. These simulation exercises aim to shed light on how significant a role remittances play in macroeconomic fluctuations and sudden stops.

Figures 3 and 4 show the ergodic distribution of bond holdings in the binding economies with and without remittance shocks for Turkey and Mexico, respectively. Those figures illustrate that remittances shocks increase

Table 2. Parameter Values

Parameter	Value	Definition	Source
General parameters			
σ	2	Relative risk aversion	RBC parametrization
y^T	1	Tradable endowment	Normalization
R	1.0159	Gross interest rate	RBC parametrization
μ	0.316	Elasticity of substitution	Ostry and Reinhart (1992)
p^N	1	Relative price of nontradables	Normalization
Country-specific parameters for Turkey			
Y^N/Y^T	1.3418	Share of nontradable output	Turkish data
Rem/GDP	0.03	Remittance-GDP ratio	Turkish data
κ	0.4	Constraint coefficient	Set to match SS dynamics
ω	0.4222	CES weight	Calibration
γ	0.0198	Elasticity of discount factor	Calibration
Country-specific parameters for Mexico			
Y^N/Y^T	1.543	Share of nontradable output	Mexican data
Rem/GDP	0.02	Remittance-GDP ratio	Mexican data
κ	0.4	Constraint coefficient	Set to match SS dynamics
Ω	0.3723	CES weight	Calibration
γ	0.0187	Elasticity of discount factor	Calibration

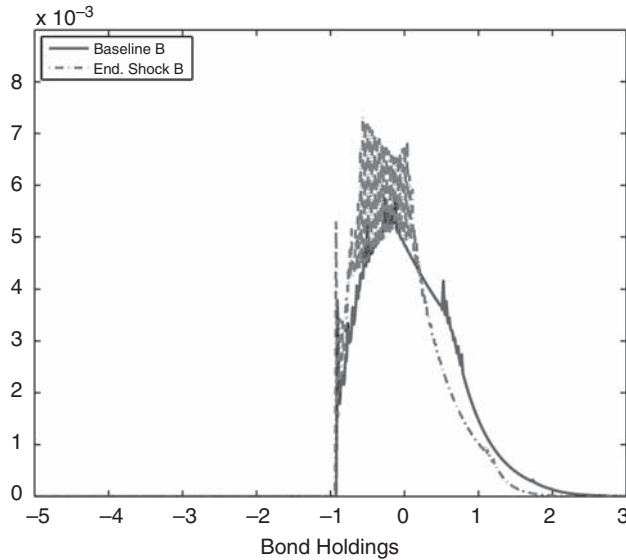
Note: This table shows the parameter values used in calibrating the model economies. The first column shows the parameters, the second column shows their values, the third column shows their definition, and the last column shows the source used in calculating those parameters. RBC=real business cycle; CES=constant elasticity of substitution.

the precautionary savings for Turkey as the distribution for the *Baseline B* is more to the right compared with the distribution for *End. Shock B*. For Mexico, however, remittances shocks decrease the precautionary savings due to their countercyclical nature. Notice that the distribution for *Baseline B* for Mexico is to the left of *End. Shock B*.

Table 3 summarizes the long-run business cycle statistics for Turkey. In the nonbinding case, the elimination of procyclical remittance fluctuations reduces the volatility of consumption from 1.77 to 1.51 percent. In line with this change, the volatility of aggregate consumption, the volatility of relative price of nontradables, and the volatility of savings also decline. Procyclical remittance fluctuations lead to stronger comovement of consumption with income (compare, for instance, the correlation of tradable consumption with GDP standing at 0.72 in the baseline case with the respective correlation of 0.67 in the endowment shock only case). Overall, remittances fluctuations reduce welfare by 0.16 percent in the nonbinding case, and 0.27 percent in the binding case.¹⁷

¹⁷Welfare calculations are performed using a compensating variation metric in consumption that equate expected lifetime utilities with and without remittance fluctuations.

Figure 3. Long-Run Distributions of Bond Holdings in the Binding Economy for Turkey



Source: Authors' calculations.

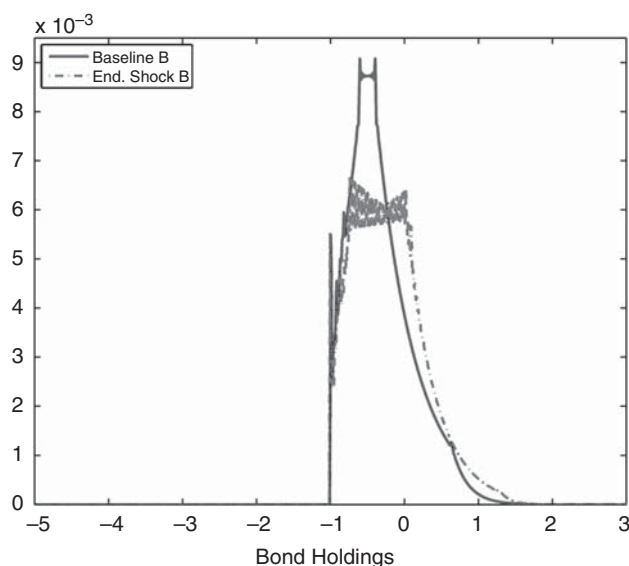
Note: The figure shows the long-run distribution of bond holdings in the binding economy (B) with and without remittances shocks. The solid line shows the results for the baseline binding economy. The dashed line shows the results for the binding economy with endowment (End.) shocks only.

The table also illustrates the strengthening effect of remittance fluctuations on precautionary savings. This strengthening effect arises because remittance fluctuations reduce the catastrophic income levels.¹⁸

Table 4 summarizes the long-run business cycle statistics for Mexico. Contrary to the Turkish case, remittance fluctuations dampen the cycles in the Mexican economy because of their countercyclical nature. In the non-binding case, volatility of consumption falls to 1.18 percent in the baseline case compared with the high of 1.40 percent in the economy with endowment shock only. In line with that result, correlation of tradable consumption with GDP increases from 0.74 in the baseline case to 0.78 in the economy with endowment shock only. Overall, the remittances fluctuations *increase* welfare

¹⁸Aiyagari (1994) shows the relationship between the catastrophic income levels and precautionary savings behavior. In his analysis, he establishes that risk averse agents have strong incentives to build up precautionary wealth to insure against the risk of state of nature in which the income stays at its lowest level forever, that is, income is at its catastrophic level. He also shows that if a structural change in the economy such as more volatile and/or more persistent income shocks reduces the catastrophic income levels, precautionary savings that agents in the economy engage in would increase (see Aiyagari (1994); Durdu (2009); and Durdu, Mendoza, and Terrones (forthcoming) for further analysis of the relationship between catastrophic income levels and precautionary savings).

Figure 4. Long-Run Distributions of Bond Holdings in the Binding Economy for Mexico



Source: Authors' calculations.

Note: The figure shows the long-run distribution of bond holdings in the binding economy (B) with and without remittances shocks. The solid line shows the results for the baseline binding economy. The dashed line shows the results for the binding economy with endowment (End.) shocks only.

by 0.0166 percent in the nonbinding economy and 0.0312 in the binding economy. As in the case for Turkey, imposition of the borrowing constraint does not change the direction of the changes as a result of remittance fluctuations.

Because emerging countries are typically borrowing constrained and the borrowing constraints in the economy get tighter on the eve of a financial crisis, an exogenous shock that can be smoothed with foreign borrowing can lead to a sudden stop in such economies. With occasionally binding borrowing constraints in place, our model can generate sudden stops, allowing us to explore how remittance fluctuations affect sudden stops by analyzing model dynamics.¹⁹ For this purpose, we use forecasting functions. To derive those functions, conditional on the economy being in a state in which the borrowing constraint binds, we first give a one-standard-deviation negative tradable endowment shock and derive the response in *End. Shock NB* economy. We then simultaneously give a one-standard-deviation negative endowment shock and a one-standard-deviation remittance shock and derive the response of the baseline *NB* economy. We choose the nature of

¹⁹Sudden stops are modeled using the same mechanism as in Mendoza (2002).

Table 3. Long-Run Business Cycle Statistics of the Model Economy Calibrated to Turkey

	Economies			
	Baseline nonbinding economy	Baseline binding economy	Endowment shock nonbinding economy	Endowment shock binding economy
Means				
Tradable consumption	1.06	1.07	1.06	1.07
Aggregate consumption	1.21	1.22	1.21	1.22
Relative price of nontradables	1.00	1.02	1.00	1.01
Savings	-0.78	0.09	-0.79	-0.07
Current account-GDP ratio	-0.03	-0.03	-0.03	-0.03
Welfare (%)	-0.16	-0.27	NA	NA
Standard deviation (%)				
Tradable consumption	1.77	1.41	1.51	1.20
Aggregate consumption	0.78	0.62	0.67	0.53
Relative price of nontradables	2.33	1.86	1.99	1.58
Savings	97.94	60.34	84.20	52.13
Current account-GDP ratio	1.48	1.43	1.48	1.45
Correlation with GDP				
Tradable consumption	0.72	0.70	0.68	0.65
Aggregate consumption	0.72	0.70	0.68	0.65
Relative price of nontradables	0.72	0.70	0.68	0.65
Savings	0.68	0.60	0.62	0.56
Current account-GDP ratio	0.74	0.77	0.80	0.82
Autocorrelation				
Tradable consumption	0.99	0.97	0.99	0.97
Aggregate consumption	0.99	0.97	0.99	0.97
Relative price of nontradables	0.99	0.97	0.99	0.97
Savings	0.99	0.99	0.99	0.99
Current account-GDP ratio	0.52	0.52	0.53	0.52

Note: The first column shows the statistics in the model economy with nonbinding borrowing constraint and with both the endowment and remittance shocks, the second column shows the statistics in the model economy with binding borrowing constraint and with both the endowment and remittance shocks. The last two columns show the statistics for the respective economies with nonbinding and binding borrowing constraints but the endowment shocks only. Welfare calculations illustrate how remittance fluctuations affect welfare compared with the “endowment shock only”-case using a compensating variation metric.

Table 4. Long-Run Business Cycle Statistics of the Model Economy Calibrated to Mexico

	Economies			
	Baseline nonbinding economies	Baseline binding economies	Endowment shock nonbinding economies	Endowment shock binding economies
Means				
Tradable consumption	1.04	1.05	1.04	1.05
Aggregate consumption	1.32	1.34	1.32	1.33
Relative price of nontradables	1.00	1.01	1.00	1.01
Savings	-0.83	-0.32	-0.83	-0.21
Current account-GDP ratio	-0.02	-0.02	-0.02	-0.02
Welfare (%)	1.66e-2	3.12e-2	NA	NA
Standard deviation (%)				
Tradable consumption	1.18	0.99	1.40	1.07
Aggregate consumption	0.47	0.40	0.56	0.43
Relative price of nontradables	1.55	1.30	1.84	1.40
Savings	63.02	42.89	78.70	48.39
Current account-GDP ratio	1.02	1.00	1.01	1.00
Correlation with GDP				
Tradable consumption	0.74	0.72	0.78	0.74
Aggregate consumption	0.74	0.73	0.78	0.74
Relative price of nontradables	0.74	0.72	0.78	0.74
Savings	0.66	0.60	0.71	0.63
Current account-GDP ratio	0.75	0.78	0.71	0.76
Autocorrelation				
Tradable consumption	0.97	0.95	0.99	0.97
Aggregate consumption	0.97	0.95	0.99	0.97
Relative price of nontradables	0.97	0.95	0.99	0.97
Savings	0.99	0.99	0.99	0.99
Current account-GDP ratio	0.60	0.60	0.60	0.60

Note: The first column shows the statistics in the model economy with nonbinding borrowing constraint and with both the endowment and remittance shocks, the second column shows the statistics in the model economy with binding borrowing constraint and with both the endowment and remittance shocks. The last two columns show the statistics for the respective economies with nonbinding and binding borrowing constraints but the endowment shocks only. Welfare calculations illustrate how remittance fluctuations affect welfare compared with the “endowment shock only”-case using a compensating variation metric.

remittance shocks by considering the opposing cyclical characteristics of remittances in Turkey and Mexico. We follow the same steps for the binding economies, as well. By taking the differences between these responses, we can calculate the additional response that remittance shocks trigger.

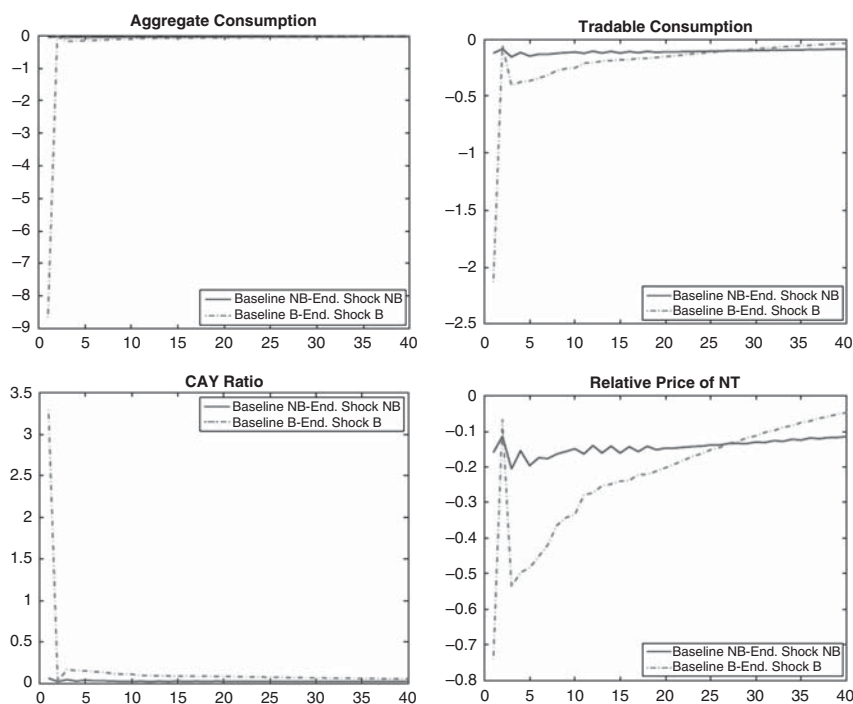
Figure 5 plots the conditional forecasting functions of aggregate consumption, tradable consumption, current account-GDP ratio, and relative price of nontradables for Turkey. In the nonbinding cases, additional (negative) remittance shocks trigger mild additional responses (notice how close solid lines in Figure 5 are to the zero line). In the binding cases, however, remittance shocks trigger much larger additional responses. For example, remittance shocks lead to around 2 percent additional decline in tradable consumption, around 8 percent additional decline in aggregate consumption, around 3 percent additional surplus in current account-GDP ratio, and 0.7 percent additional decline in relative price of nontradables. These results suggest that on the eve of a financial crisis, remittance shocks can have a significant effect on the economy, as the borrowing constraints in the economy get tighter and small additional shocks get magnified by a Fisherian debt deflation process (see Mendoza (2005) for further analysis of Fisherian debt deflation).

Figure 6 plots the conditional forecasting function for Mexico. In this exercise, we compare the effect of a negative endowment shock alone with that of a negative endowment shock together with a positive remittance shock by considering the countercyclical nature of remittances in Mexico. In line with the results for the Turkish case, when the economy is not borrowing constrained, the remittance shocks do not alter the responses in the economy significantly. When the economy is borrowing constrained, however, positive remittance shocks provide a significant smoothing effect. For example, the positive remittance shock helps reduce the decline in tradable consumption by about 1.4 percent, that in aggregate consumption and relative price of nontradables by around 6 percent and 0.5 percent, respectively, while smoothing current account reversal by about 2 percentage points.

III. Conclusion

Recently presented evidence indicates that remittances sent home by Turkish workers abroad move in the same direction as the business cycles in Turkey, whereas remittance receipts of Mexico are countercyclical. Given the procyclicality of remittances received by Turkey, drops in the amounts remitted by migrant Turkish workers during or shortly after cyclical contractions in Turkish GDP will tend to fan the flames of a crisis in the Turkish economy, whereas countercyclical remittances from migrant Mexican workers will tone down a crisis in the Mexican economy. Yet, the existing literature is largely silent about the magnitude of the effects of remittances on cyclical volatility of output and other macroeconomic variables, as well as sudden stops, particularly in the case of countries whose remittance receipts fluctuate procyclically.

Figure 5. Conditional Forecasting Functions in the Model Economy Calibrated to Turkey

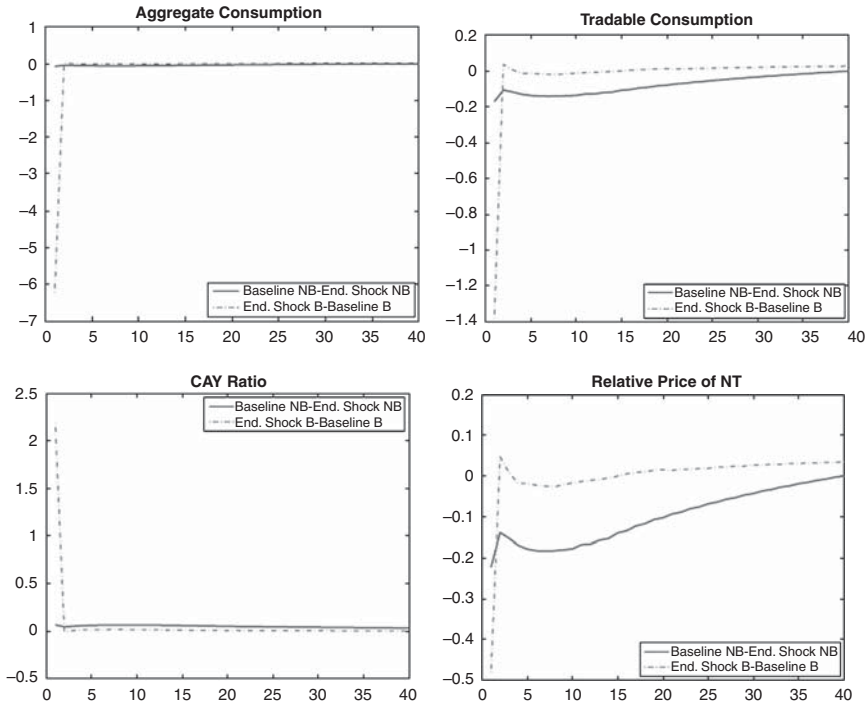


Source: Authors' calculations.

Note: The figure shows the conditional forecasting functions for aggregate consumption, tradable consumption, current account-GDP (CAY) ratio, and relative price of nontradables (NT). The solid line shows the relative response in the nonbinding economy (NB) with both remittance and endowment (End.) shocks relative to the nonbinding economy with endowment shock only. The dashed line shows relative response in the binding economy (B) with both remittances and endowment shocks relative to the binding economy with endowment shock only.

This paper aimed to close this gap in the literature by exploring the effects of migrants' remittances flows with opposite responses to business cycle fluctuations in the recipient economies on key macroeconomic aggregates and sudden stops experienced by these countries. For this purpose, we considered Mexico and Turkey, two emerging economies that rank among the major recipients of remittances, whose receipts are countercyclical and procyclical to home business cycles, respectively. We employed a general equilibrium framework with a tradable and a nontradable sector to model small open economies of Mexico and Turkey. We allowed for dollarization of the liabilities and let agents face a borrowing constraint in international capital markets for added realism in capturing the common structural characteristics of Mexican and Turkish economies. After calibrating this model to the data for each economy, we ran four different simulation experiments involving an endowment and a remittance shock for each

Figure 6. Conditional Forecasting Functions in the Model Economy Calibrated to Mexico



Source: Authors' calculations.

Note: The figure shows the conditional forecasting functions for aggregate consumption, tradable consumption, current account-GDP (CAY) ratio, and relative price of nontradables (NT). The solid line shows the relative response in the nonbinding economy (NB) with both remittance and endowment (End.) shocks relative to the nonbinding economy with endowment shock only. The dashed line shows relative response in the binding economy (B) with both remittances and endowment shocks relative to the binding economy with endowment shock only.

country, under binding and nonbinding borrowing constraint scenarios. Our results indicated the following:

- Remittances dampen the cycles in Mexico, whereas they amplify the cycles in Turkey.
- Long-run effects of remittances do not significantly depend on the existence of borrowing constraints, but their short-run effects depend on whether the economy is borrowing constrained or not.
- Fisherian debt-deflation can magnify the effects of fluctuations in remittances in the short run in both countries.
- Countercyclical (or procyclical) remittance fluctuations can help to reduce precautionary savings by increasing (or reducing) catastrophic income levels.

- On the eve of a financial crisis, remittances packages received from abroad could significantly reduce (or increase) the impact effect of financial crises if the remittances are countercyclical (or procyclical), implying that it could indeed pour, when it rains in the case of procyclical remittances (as Kaminsky, Reinhart, and Végh put it in their 2004 study on procyclical capital flows).

These findings provide a rationale for the importance of remittances in mitigating macroeconomic fluctuations and sudden stops. While helping to close an important gap in the literature, our study did not explore the migrants' decision on how much and when to remit, that is, what makes remittances procyclical or countercyclical. We rather took those cyclical properties of remittance fluctuations as given and looked at their implications. To derive these properties, we would need to endogenize the altruistic motives for remittance decisions. Albeit interesting, this task is left for further research.

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